

## Introduction

Penalized regression on B-splines (P-spline[1]), is a flexible and stable approach for fitting smooth functional effects in structured additive regression models[2].

$$\mathbf{y} = f(\mathbf{x}) + \boldsymbol{\epsilon}; \quad \boldsymbol{\epsilon} \sim N(0, \tau_{\boldsymbol{\epsilon}}^{-1} \mathbf{I})$$

$$f(\mathbf{x}) = \mathbf{B}\boldsymbol{\beta}$$

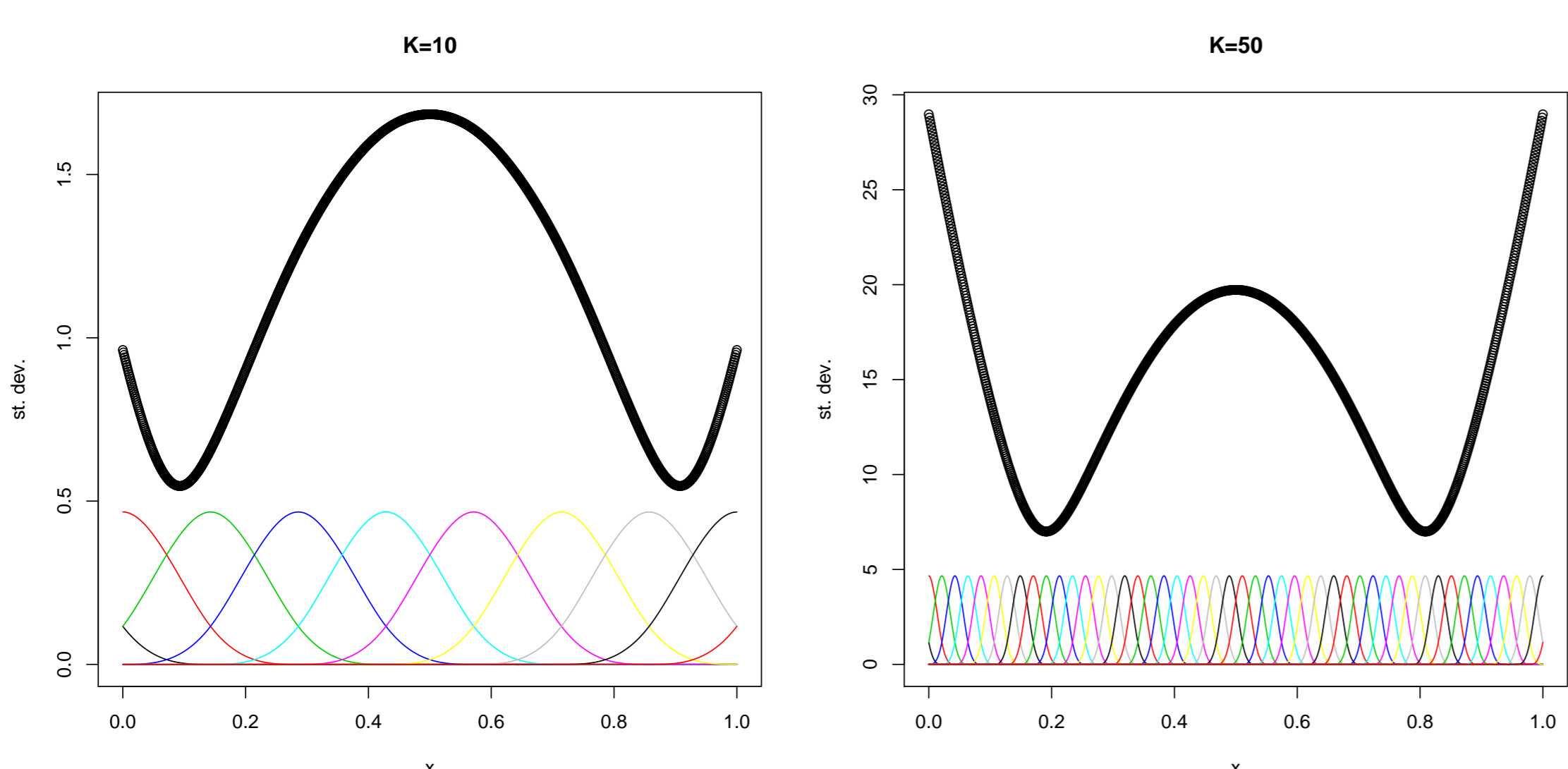
$\mathbf{B}$  is an  $n \times K$  matrix of B-splines (equally-spaced over the covariate domain) and  $\boldsymbol{\beta}$  a vector of coefficients.

### Bayesian P-splines[3]: IGMRF prior

$$\pi(\boldsymbol{\beta}|\tau_{\boldsymbol{\beta}}) = (2\pi)^{-\text{rank}(\mathbf{R})/2} (|\tau_{\boldsymbol{\beta}} \mathbf{R}|^*)^{1/2} \exp\left\{-\frac{\tau_{\boldsymbol{\beta}}}{2} \boldsymbol{\beta}^T \mathbf{R} \boldsymbol{\beta}\right\}$$

$$\pi(\tau_{\boldsymbol{\beta}}) = \text{Gamma}(a, b)$$

The IGMRF  $\pi(\boldsymbol{\beta}|\tau_{\boldsymbol{\beta}})$  models deviation from a polynomial of degree  $K - \text{rank}(\mathbf{R}) - 1$ ; see [4]. How large the deviation depends on  $\tau_{\boldsymbol{\beta}}$ . Elicitation of a prior for  $\pi(\tau_{\boldsymbol{\beta}})$  is critical; see the scaling issue in Fig 1.



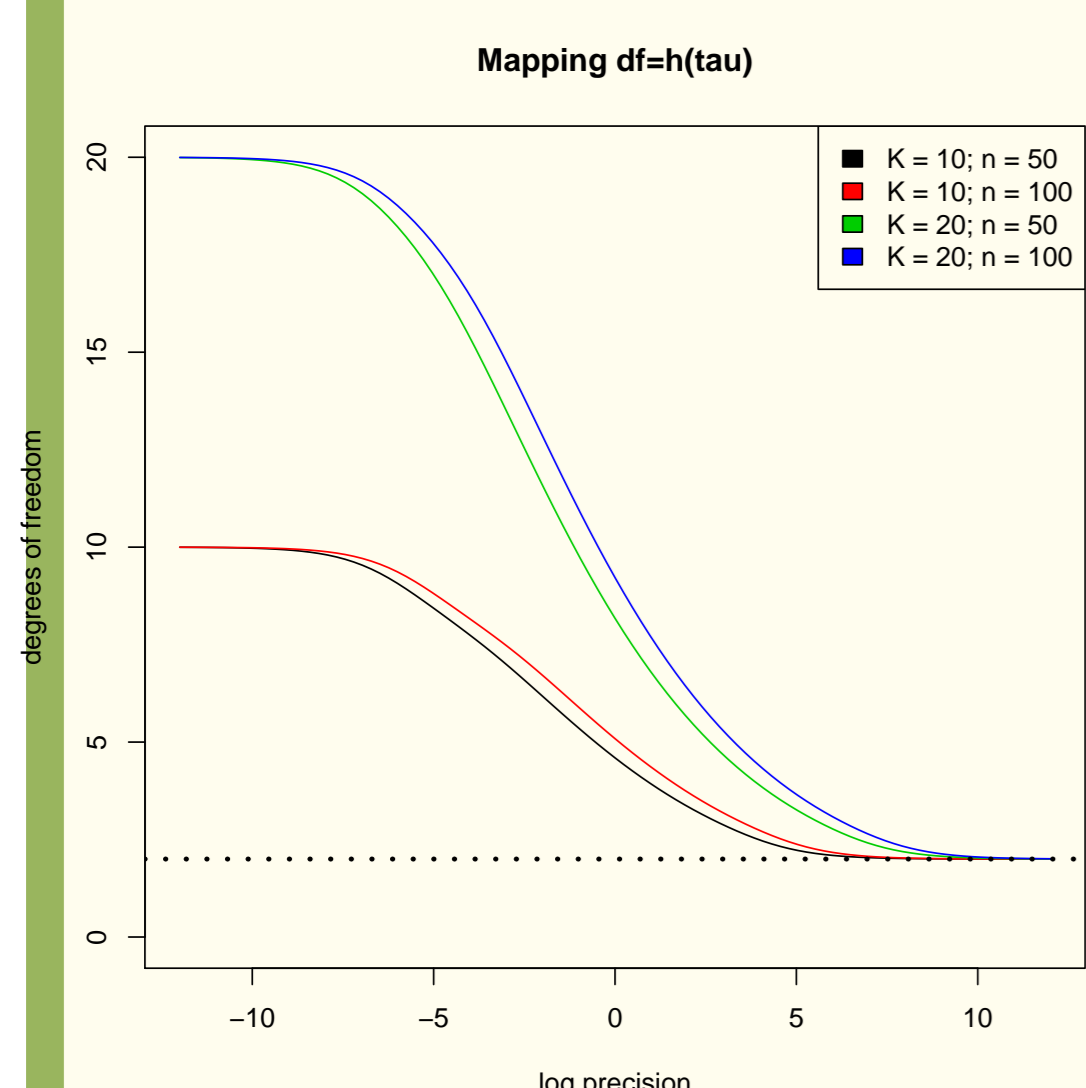
**Fig 1: Scaling issue[5].** The marginal standard deviation of the IGMRF,  $\mathbf{B}(\tau_{\boldsymbol{\beta}} \mathbf{R})^{-1} \mathbf{B}^T$ , scales with design (i.e., it changes with the number of knots, i.e. with  $K$ ).

## Main idea

To overcome the scaling issue we consider building priors on an interpretable property of the model, indicating the complexity of the smooth function  $f$ , e.g., the **number of effective degrees of freedom**[6].

$$df = \text{tr} \left\{ \left( \mathbf{B}^T \mathbf{B} + \frac{\tau_{\boldsymbol{\beta}}}{\tau_{\boldsymbol{\epsilon}}} \mathbf{R} \right)^{-1} \mathbf{B}^T \mathbf{B} \right\}$$

$$= \text{tr} \left\{ \left( \mathbf{I} + \frac{\tau_{\boldsymbol{\beta}}}{\tau_{\boldsymbol{\epsilon}}} \mathbf{R} (\mathbf{B}^T \mathbf{B})^{-1} \right)^{-1} \right\} = \sum_{k=1}^K \frac{1}{1 + \frac{\tau_{\boldsymbol{\beta}}}{\tau_{\boldsymbol{\epsilon}}} v_k}$$

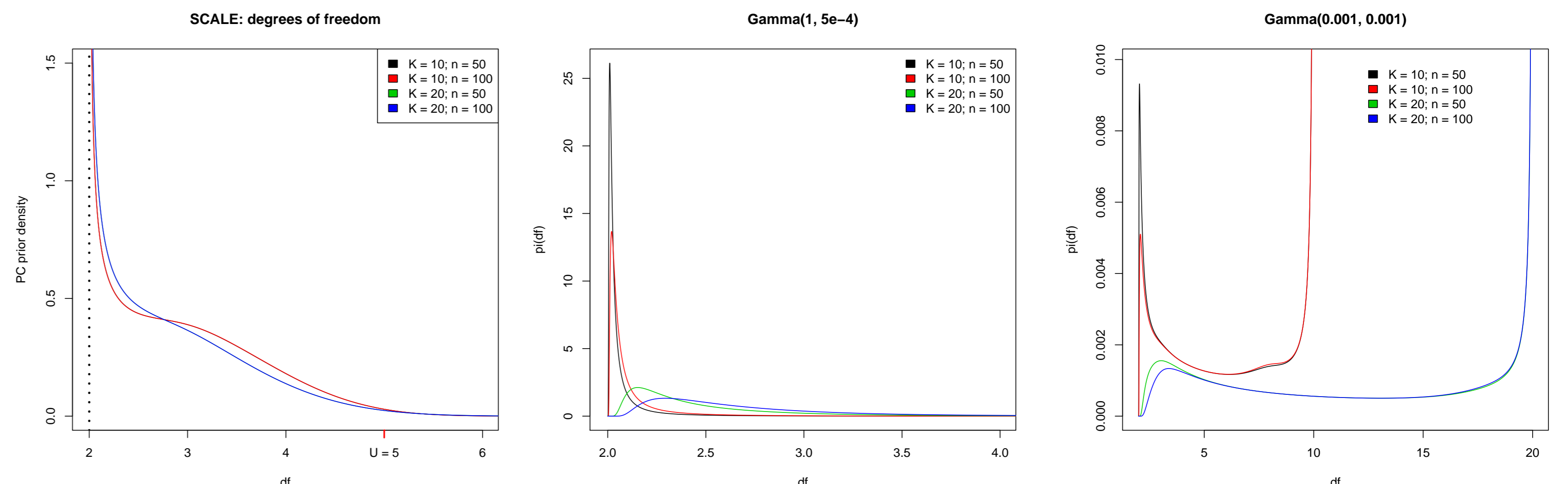


$df \in (K - \text{rank}(\mathbf{R}), K)$ ; it is a function of  $\tau_{\boldsymbol{\beta}}$ , given  $\tau_{\boldsymbol{\epsilon}}$  and relates to the **degree of an equivalent polynomial** (equivalent to the smooth function  $\mathbf{B}\boldsymbol{\beta}$ );  $v = \text{eigen}(\mathbf{R}(\mathbf{B}^T \mathbf{B})^{-1})$  val

## Penalized Complexity (PC) priors

The four principles of PC priors[7] applied to P-splines

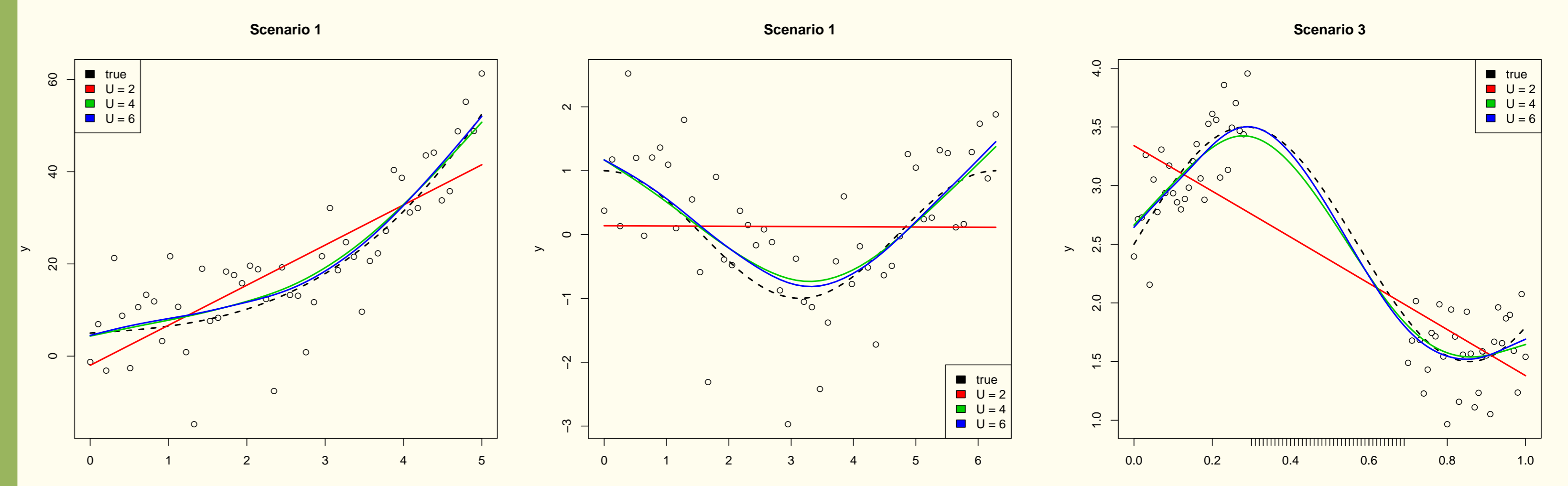
- 1) Parsimony:** the *base model* is the best unless there is enough evidence for more a *flexible model*. We denote with  $f_0 = \mathbf{B}\boldsymbol{\beta}_0$  the base model and with  $f = \mathbf{B}\boldsymbol{\beta}$  the flexible one. Model  $f_0$  is a polynomial of degree equal to the rank deficiency of  $\mathbf{R}$  minus 1.
- 2) Measure of complexity:** the *distance* between two competing models is measured by  $d = \sqrt{2\text{KLD}}$ ; KLD is the Kullback Leibler Divergence. From [7]: for  $\tau_{\boldsymbol{\beta}} \ll \tau_{\boldsymbol{\beta}_0}$  and  $\tau_{\boldsymbol{\beta}_0} \rightarrow \infty$ , the quantity  $\text{KLD}(\boldsymbol{\beta}||\boldsymbol{\beta}_0)$  goes to  $\frac{\tau_{\boldsymbol{\beta}_0} K}{2\tau_{\boldsymbol{\beta}}}$ , then  $d = \sqrt{2\text{KLD}(\boldsymbol{\beta}||\boldsymbol{\beta}_0)} = \sqrt{\tau_{\boldsymbol{\beta}_0} K / \tau_{\boldsymbol{\beta}}}$ .
- 3) Constant rate penalization:** for increasing  $d$ , flexible models are penalized by a constant decay rate. The PC prior is an exponential distribution on the distance scale,  $\pi_{PC}(d) = \lambda \exp(-\lambda d)$ , with rate  $\lambda$ . By a change of variable and setting the rate  $\lambda = \theta / \sqrt{K\tau_{\boldsymbol{\beta}_0}}$ , [7] obtain the PC prior for  $\tau_{\boldsymbol{\beta}}$  as a Gumbel(1/2,  $\theta$ ) type 2 distribution.
- 4) User-defined scaling:** define  $\theta$  by eliciting an upper bound for the degree of an equivalent polynomial for  $f$ . Let us denote the mapping between  $\tau_{\boldsymbol{\beta}}$  and  $df$  (conditional on  $\tau_{\boldsymbol{\epsilon}}$ ) as  $h(\tau_{\boldsymbol{\beta}}|\tau_{\boldsymbol{\epsilon}}) = \sum_{k=1}^K \frac{1}{1 + \frac{\tau_{\boldsymbol{\beta}}}{\tau_{\boldsymbol{\epsilon}}} v_k}$ . Let  $U$  be an upper bound for  $df$  and  $\alpha$  the tail event probability, we need  $\theta$  such that  $\alpha = \text{Pr}(h(\tau_{\boldsymbol{\beta}}) > U) = F(h^{-1}(U))$ , where  $F$  is the Gumbel c.d.f, PC prior conditional on  $\tau_{\boldsymbol{\epsilon}}$ :  $\pi_{PC}(df|U, \alpha) = \pi_{PC}(\tau_{\boldsymbol{\beta}}|\theta = -\log(\alpha)\sqrt{(h^{-1}(U))})$



**Fig 2: Implied degrees of freedom** for the PC prior  $\pi_{PC}(U = 5, \alpha = 0.01)$  (left), the Gamma(1, 5e-4) and Gamma(0.001, 0.001);  $\tau_{\boldsymbol{\epsilon}}$  set equal to 1.

## P-splines with joint prior $\pi(\tau_{\boldsymbol{\beta}}, \tau_{\boldsymbol{\epsilon}}) = \pi(\tau_{\boldsymbol{\beta}}|\tau_{\boldsymbol{\epsilon}})\pi(\tau_{\boldsymbol{\epsilon}})$

$$\pi(\tau_{\boldsymbol{\beta}}|\tau_{\boldsymbol{\epsilon}}) = \text{Gumbel}(1/2, \theta(\tau_{\boldsymbol{\epsilon}})) \quad \pi(\tau_{\boldsymbol{\epsilon}}) \propto 1/\tau_{\boldsymbol{\epsilon}} \quad (1)$$



**Fig 3: The joint prior in action**, in three examples with simulated data.

## Concluding remarks

- The PC prior  $\pi_{PC}(df|U, \alpha)$  does not force overfitting (non zero probability at the base model) and has the same interpretation under different designs.
- To fit P-splines with joint prior (1) we use MCMC:  $\theta(\tau_{\boldsymbol{\epsilon}})$  has to be recomputed at each iteration, conditional on the current  $\tau_{\boldsymbol{\epsilon}}$ .
- Future work will include generalization to non Gaussian data.